

1                   **Natural variability has slowed the decline in**  
2                   **western-US snowpack since the 1980s**

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8                   **Key Points:**

- 9                   • Circulation changes have offset most western-US snowpack loss from global warm-  
10                  ing since the 1980s.
- 11                  • Circulation changes are likely a result of natural variability, not anthropogenic forc-  
12                  ing.
- 13                  • Snowpack loss will likely accelerate in coming decades as the phase of natural vari-  
14                  ability subsides.

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**Abstract**

Spring snowpack in the mountains of the western United States (US) has not declined substantially since the 1980s, despite significant global and regional warming. Here we show that this apparent insensitivity of snowpack to warming is a result of changes in the atmospheric circulation over the western US, which have reduced snowpack losses due to warming. Climate model simulations indicate that the observed circulation changes have been driven in part by a shift in Pacific sea-surface temperatures that is attributable to natural variability, and not part of the simulated response to anthropogenic forcing. Removing the influence of natural variability reveals a robust anthropogenic decline in western-US snowpack since the 1980s, particularly during the early months of the accumulation season (October-November). These results suggest that the recent stability of western-US snowpack will be followed by a period of accelerated decline once the current mode of natural variability subsides.

**Plain language summary**

Melting snowpack is a vital source of water in the western US during the summer, when rainfall is usually scarce. Although the amount of water contained in the snowpack has declined over the past century, it has been surprisingly stable since the 1980s, despite 1°C of warming over the same period. At first glance, this result might appear to indicate that the snowpack is quite resilient to warming. However, here we show that the contribution of global warming to western-US snowpack loss has in reality been large and widespread since the 1980s, but mostly offset by natural variability in the climate system. This result points to a faster rate of snowpack loss in coming decades, when the impact of global warming is more likely to be amplified, rather than offset, by natural variability.

**1 Introduction**

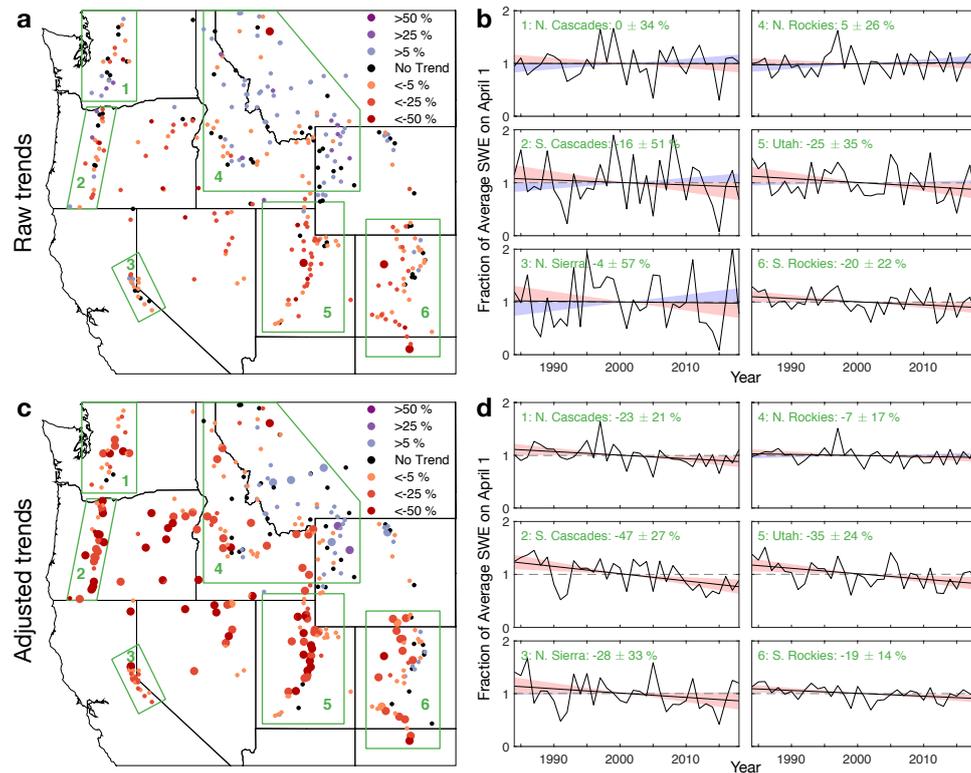
Snowpack in the western United States (US) acts as a natural reservoir, storing water during the cool wet season (October - March) and releasing it to the landscape during the warm dry season (Cayan, 1996; Barnett et al., 2005). Containing more than 150 cubic kilometers of water at its peak—usually reached around April 1—the western-US winter snowpack holds about as much water as all the man-made reservoirs in the western US combined (Mote et al., 2018). Since 1950, the snowpack on April 1 has decreased by 15-30% over much of the western US, as warmer temperatures have caused a shift from

47 snow to rain, particularly at low elevations (Mote et al., 2018). Models indicate a fur-  
48 ther decrease in winter snowpack of perhaps 60% by 2050 (Ashfaq et al., 2013; Fyfe et  
49 al., 2017), leading to a dramatic reduction in summer stream flows in a region where wa-  
50 ter shortages are already common (Christensen et al., 2004; Barnett et al., 2005)

51 Despite this alarming forecast, snowpack has been surprisingly stable in recent decades.  
52 Fig. 1a shows the trends in April 1 snowpack at 329 automated Snowfall Telemetry (SNO-  
53 TEL) stations where the average April 1 snowpack contains at least 25 cm of water-equivalent.  
54 Trends were calculated over 35 winters beginning in 1983-84, when the network reached  
55 nearly its current level of coverage (Fig. S1), and one year after the record-setting El Niño  
56 of 1982-83. During this 35-year period, only four sites experienced a statistically signif-  
57 icant decline in April 1 snowpack (95% confidence), while 208 sites (63%) experienced  
58 an insignificant decline. The other 117 sites (36%) experienced a positive (but statisti-  
59 cally insignificant) trend. We find a similar result when we repeat the analysis using regionally-  
60 averaged time series of April 1 snowpack (Fig. 1b), with no region exhibiting a statis-  
61 tically significant trend since 1983-84. This result is consistent with Mudryk et al. (2018),  
62 who found similar stability in the spring snowpack of western Canada over roughly the  
63 same time period. Meanwhile, the average winter surface temperature over the western  
64 US increased by more than 1 °C during this period—on par with the average warming  
65 trend across global land surfaces (Fig. S2). Taken at face value, this result suggests that  
66 winter snowpack in the western US may be remarkably insensitive to warming, seem-  
67 ingly contradicting model forecasts of rapid decline (Ashfaq et al., 2013; Fyfe et al., 2017).

76 On decadal timescales, however, snowpack is influenced not only by anthropogenic  
77 warming, but also by natural variability in the large-scale atmospheric circulation (Deser  
78 et al., 2012). Depending on its phase, natural variability can either offset or accelerate  
79 the changes in snowpack due to anthropogenic warming (Cayan, 1996; Mote, 2006; Stoelinga  
80 et al., 2010; McCabe & Wolock, 2009; Smoliak et al., 2010). In order to quantify the an-  
81 thropogenic trend in western-US snowpack, one must first remove the component of the  
82 trend that can be attributed to natural variability in the atmosphere.

83 One approach that previous studies have taken to minimize the influence of nat-  
84 ural variability has been to focus on the ratio of snowpack to accumulated precipitation  
85 (S/P), which tends to be more sensitive than snowpack to changes in temperature (Barnett  
86 et al., 2008; Pierce et al., 2008; Pierce & Cayan, 2013). Consistent with these studies,



68 **Figure 1.** Raw and dynamically adjusted trends in April 1 snowpack at 329 SNOTEL sites  
 69 across the western US between 1984 and 2018 (water years). a) Raw trends. Large circles are  
 70 statistically significant ( $p = 0.05$ ). b) Regionally-averaged time series and trend estimates for six  
 71 regions outlined in (a): 1: northern Cascades; 2: southern Cascades; 3: northern Sierra Nevada;  
 72 4: northern Rocky Mountains; 5: Utah; 6: southern Rocky Mountains. Shaded cones represent  
 73 the 95% confidence range of the trend estimate, with red indicating a negative trend, and blue  
 74 indicating a positive trend. c,d) As in (a,b), but after dynamical adjustment. All values are  
 75 expressed as a fraction of the 35-year average April 1 snowpack.

87 we too find that SNOTEL stations exhibit more robust declines in S/P than in snow-  
 88 pack alone (Fig. S3 vs. Fig. 1a,b). However, while S/P may exhibit a clearer warming  
 89 signal, it provides little insight into the magnitude of snowpack changes resulting from  
 90 anthropogenic warming versus natural variability.

91 In this paper, we use a method called “dynamical adjustment” to quantify the in-  
 92 fluences of both anthropogenic warming and natural variability on April 1 snowpack at  
 93 each of the 329 SNOTEL stations shown in Fig. 1a. We find that, since 1983-84, changes  
 94 in the atmospheric circulation have contributed to a  $\sim 30\%$  increase in April 1 snow-

95 pack in the Cascade Mountains and northern Sierra Nevada, and an increase of  $\sim 10\%$   
96 in Utah and the Northern Rocky Mountains, offsetting much of the decline due to global  
97 warming. Simulations performed with an atmospheric general circulation model (GCM)  
98 and prescribed historical sea-surface temperatures (SSTs) indicate that the observed cir-  
99 culation change has likely been driven in part by internal atmospheric variability, and  
100 in part by a shift in Pacific SSTs toward the cool phase of the Interdecadal Pacific Os-  
101 cillation. We find that such SST/circulation trends are not part of the simulated response  
102 to anthropogenic forcing in coupled GCMs, but are likely associated with low-frequency  
103 natural variability that will eventually subside, ushering in a period of accelerated snow-  
104 pack loss.

## 105 2 Data and Methods

### 106 2.1 Dynamical Adjustment

107 While there are various ways to perform dynamical adjustment (Deser et al., 2016;  
108 Lehner et al., 2017; Merrifield et al., 2017; O’Reilly et al., 2017; Fereday et al., 2018; Lehner  
109 et al., 2018), here we use the well-known method of partial least-squares (PLS) regres-  
110 sion (Smoliak et al., 2010; Wallace et al., 2012; Deser et al., 2014; Smoliak et al., 2015;  
111 Christian et al., 2016). At each SNOTEL station, we adjust the 35-year time series of  
112 net snowpack accumulation (i.e., the “predictand”) during three separate stages of the  
113 accumulation season: early (October-November), middle (December-January), and late  
114 (February-March). Net accumulation during each stage was computed from daily obser-  
115 vations of snow-water equivalent (SWE) supplied by the Natural Resources Conserva-  
116 tion Service. As predictors, we use three variables from MERRA-2 reanalysis (Gelaro  
117 et al., 2017) that are representative of the large-scale atmospheric circulation: (1) pres-  
118 sure at mean sea level, and geopotential height at (2) 500 hPa and (3) 250 hPa. We re-  
119 move the global means of (2) and (3) to account for global warming. We restrict the do-  
120 main of each predictor field to the Northern Hemisphere south of 60 N, and to the Pa-  
121 cific/North America sector between 110 and 290 E. However, our results are essentially  
122 unchanged when the domain is expanded to the full Northern Hemisphere (not shown).

123 To begin, we apply a 15-year highpass filter to both the predictand and predictors,  
124 and compute the correlation matrix  $\mathbf{X}$  between the predictand and predictors at each  
125 grid point. Filtering avoids fitting to the anthropogenic trend, and/or to low-frequency

126 variability that could impact the trend on decadal timescales. We then project the ob-  
 127 served predictor anomalies onto  $\mathbf{X}$ , producing a time series  $\mathbf{y}$ . The relationship between  
 128  $\mathbf{X}$  and  $\mathbf{y}$  is directly analogous to the relationship between an empirical orthogonal func-  
 129 tion and its corresponding principal component (Siler et al., 2018). Finally, we regress  
 130  $\mathbf{y}$  out of both the predictand and predictors, and repeat the entire procedure two more  
 131 times to yield the dynamically-adjusted snowpack accumulation. The adjusted April 1  
 132 snowpack is found by adding the adjusted sub-seasonal accumulations. While the full  
 133 set of predictor patterns is unwieldy, Fig. S4 gives an example of the average predictor  
 134 patterns among SNOTEL stations in the southern Cascades of Oregon—a region where  
 135 changes in the atmospheric circulation have had a particularly large influence on recent  
 136 snowpack trends, as discussed in Section 3.

137 To minimize overfitting, we use leave-p-out cross-validation, so that each year’s ad-  
 138 justment is based on the correlation matrix from all other years. To evaluate the suc-  
 139 cess of this approach, we have tested the same algorithm on thousands of sets of 10 randomly-  
 140 generated sub-seasonal time series, designed to approximate the  $\sim 10$  degrees of free-  
 141 dom in western-US snowpack. Using the same sub-seasonal predictor fields described above,  
 142 we find that dynamical adjustment reduces the average season-total variance by  $5.1 \pm$   
 143  $5.5\%$  ( $1\sigma$ ). By comparison, the average variance in April 1 snowpack across all SNOTEL  
 144 stations decreases by 47.2% after dynamical adjustment. This demonstrates that the large  
 145 majority of snowpack variance removed through dynamical adjustment is physically re-  
 146 lated to variability in the atmospheric circulation, and not a result of overfitting.

## 147 2.2 Trend Significance

148 Snowpack trends are calculated using linear least-squares regression. To determine  
 149 trend significance (both raw and adjusted), we first compute the  $p$  value of the trend in  
 150 each time series as if it were a single hypothesis test, accounting for year-to-year persis-  
 151 tence (Wilks, 2011). Then, to account for multiple hypotheses, we apply the false-discovery-  
 152 rate test to this set of  $p$  values, choosing a two-tailed confidence level of 95% ( $\alpha = 0.05$ ),  
 153 and adjusting for fewer degrees of freedom due to strong spatial correlations (Wilks, 2016).

### 2.3 Other Observations

Gridded sea-surface temperature (SST) observations are taken from version 2 of the monthly NOAA Optimal Interpolation SST data set (OISSTv2) (Reynolds et al., 2002). Global-mean and regional-mean surface warming is estimated from the gridded Goddard Institute for Space Studies Surface Temperature (GISTEMP) analysis (Hansen et al., 2010).

### 2.4 Model Simulations

Prescribed-SST simulations are provided by NOAA’s Facility for Climate Assessments (FACTS). We analyze 30 simulations performed with the ECHAM5 atmospheric general circulation model (GCM) at  $0.75 \times 0.75^\circ$  resolution with prescribed historical SSTs (Roeckner et al., 2003). As of October, 2018, ECHAM5 was the only GCM with a large ensemble of prescribed-SST simulations spanning the entire 35-year analysis period. Coupled GCM simulations were performed as part of the latest Coupled Model Intercomparison Project (CMIP5) (Taylor et al., 2012). We have included all available models and ensemble members in our analysis. Trends were calculated for each ensemble member from 1983-84 through 2017-18 by concatenating output from historical (through 2005) and RCP8.5 (after 2005) experiments, which best approximate historical emissions (Hayhoe et al., 2017).

## 3 Results

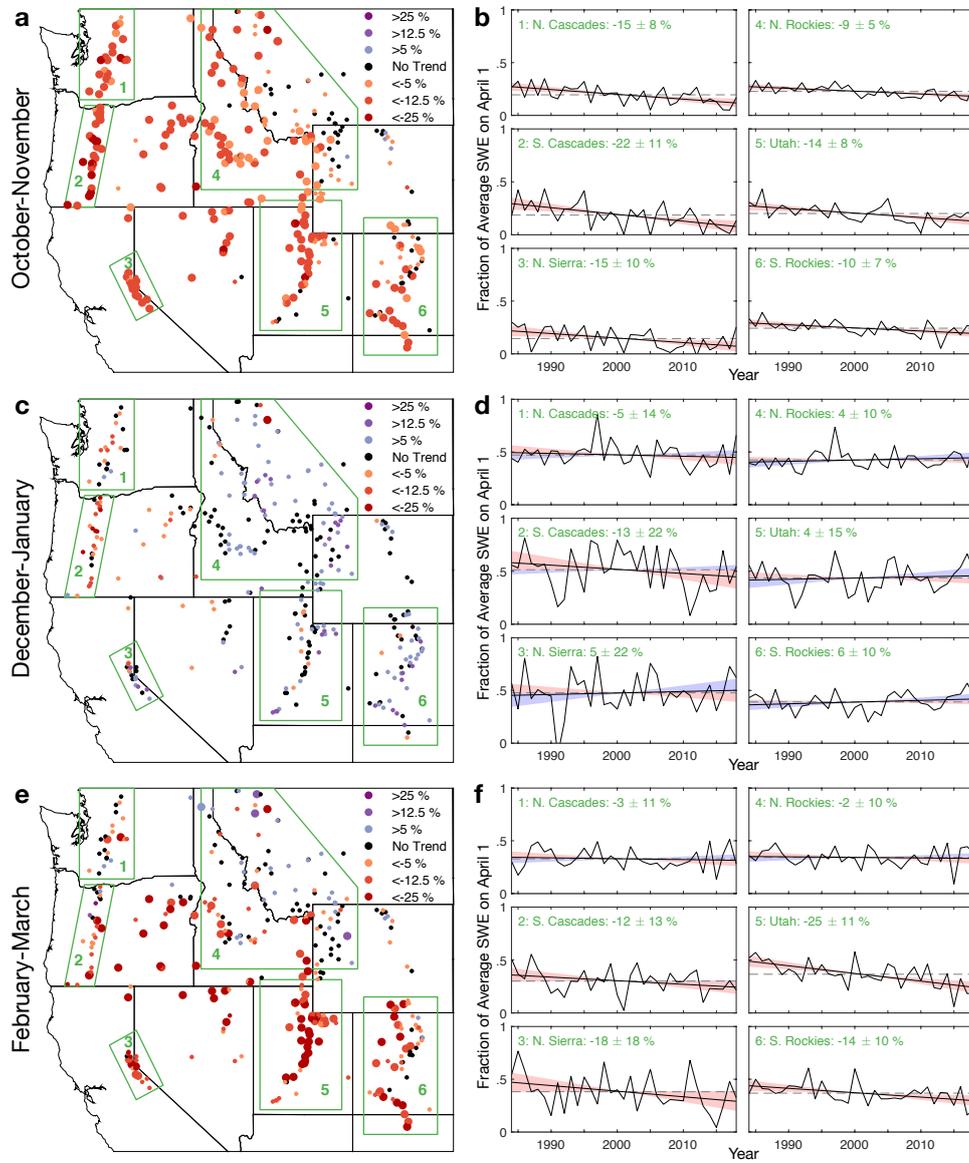
Fig. 1c shows the trends in April 1 snowpack between 1983-84 and 2017-18 after dynamically-adjusting the net sub-seasonal accumulation. In contrast to the raw time series (Fig. 1a), which exhibit almost no significant changes since 1984, the trends in the dynamically-adjusted time series are significantly negative at 106 sites (32%), and significantly positive at just 7 sites (2%), all in northwestern Wyoming or southwestern Montana. When applied to the regionally-averaged time series, dynamical adjustment reveals significant declines in every region except the northern Sierra Nevada and the northern Rocky Mountains, where the trends are negative but not statistically significant (Fig. 1d).

Comparing the three stages of the accumulation season (Fig. 2), we find that the most robust changes have occurred early in the season, when 199 sites (60%) show a sig-

184 nificant decline (Fig. 2a-b). Early-season declines are largest on the West Coast and in  
185 Utah, where above-freezing mean temperatures leave the snowpack particularly vulner-  
186 able to warming (Fig. S5a). Late season declines are less widespread (Fig. 2e-f), but still  
187 significant at 94 sites (29%), including most of Utah and the southern Rockies (regions  
188 5 and 6). In contrast, mid-season snowpack is less sensitive to warming (Fig. 2c-d), likely  
189 because the average temperature in December and January has remained below freez-  
190 ing at most sites (Fig. S5b) (Kapnick & Hall, 2012).

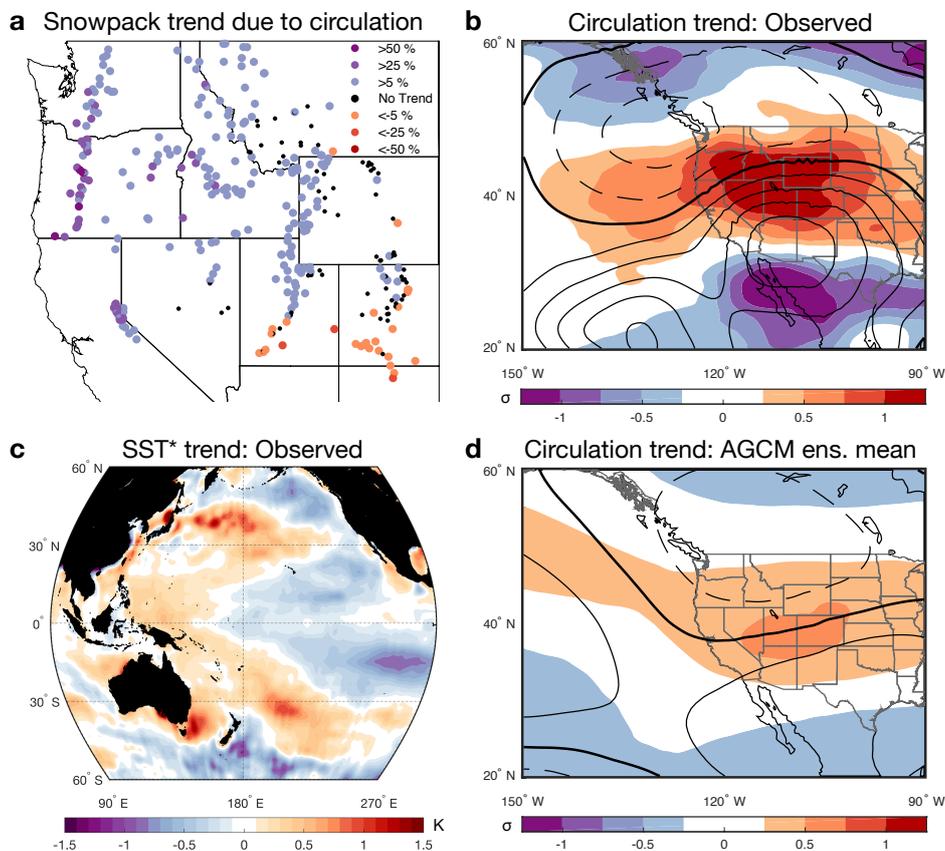
195 The increased statistical significance of the dynamically-adjusted trends compared  
196 to the raw trends can be explained in part by the smaller variance of the adjusted time  
197 series after the circulation-driven variability is removed. But the effect of dynamical ad-  
198 justment goes beyond a simple reduction in variance; it also includes substantial changes  
199 in the trends themselves. In the southern Cascades of Oregon (region 2), for example,  
200 the raw time series exhibits a decrease in April 1 snowpack of  $16 \pm 51\%$ , while the ad-  
201 justed time series exhibits a much larger decrease of  $47 \pm 27\%$  (all percentage changes  
202 are relative to the 1984-2018 mean April 1 snowpack, while uncertainty estimates indi-  
203 cate 95% confidence). The difference—an increase of 31%—represents the change in snow-  
204 pack due to changes in the atmospheric circulation. Similarly positive contributions to  
205 snowpack trends from the atmospheric circulation are found over most of the western  
206 US, with every region except the southern Rocky Mountains exhibiting adjusted trends  
207 that are at least 10% more negative than the raw trends in the same region. These dif-  
208 ferences imply that, in the absence of changes in the atmospheric circulation, the decline  
209 in western-US snowpack since the 1980s would have been substantially larger than has  
210 been observed.

211 To better understand the influence of the atmospheric circulation on western-US  
212 snowpack trends, we show the dynamical contribution to snowpack trends at each SNO-  
213 TEL site (Fig. 3a) alongside the trends in the winter-mean circulation over the same time  
214 period (Fig. 3b). The dynamical contribution in Fig. 3a is equal to the difference be-  
215 tween the raw and adjusted snowpack trends in Figs. 1a and 1c. It averages 13% across  
216 all SNOTEL sites, but is substantially larger in the Cascades and Sierra Nevada (regions  
217 1-3), where it averages 29%. These dynamically-induced snowpack trends have coincided  
218 with enhanced zonal winds at 500 hPa ( $U_{500}$ ; Fig. 3b, colors), driven by a steepening of  
219 the north-south gradient in geopotential height ( $Z_{500}$ ) over the western US (Fig. 3b, black  
220 contours). Stronger  $U_{500}$  is associated with greater moisture transport to the West Coast,



191 **Figure 2.** Dynamically-adjusted trends in snowpack accumulation during the early (October-  
 192 November; top), middle (December-January), and late (February-March; bottom) stages of the  
 193 accumulation season. Like in Fig. 1, all values are expressed as a fraction of the 35-year average  
 194 in April 1 snowpack.

221 and with larger orographic enhancement of precipitation at high elevations (Luce et al.,  
 222 2013; Siler et al., 2013), consistent with the positive dynamical contribution to snowpack  
 223 trends in Fig. 3a.



224 **Figure 3.** a) The dynamical contribution to April 1 snowpack trends between 1984 and 2018,  
 225 equivalent to the difference between the raw and adjusted trends in Figs. 1a and 1c). b) The  
 226 observed trend in winter-mean  $U_{500}$  (colors) and  $Z_{500}$  (contours), normalized by the inter-annual  
 227 standard deviation ( $\sigma$ ) at each grid point. c) The observed trend in winter-mean SST\*, defined as  
 228 the local SST minus the average SST within the region shown. d) As in (b), but for the ensemble  
 229 mean of 30 simulations from the ECHAM5 atmospheric GCM run with prescribed historical  
 230 SSTs.

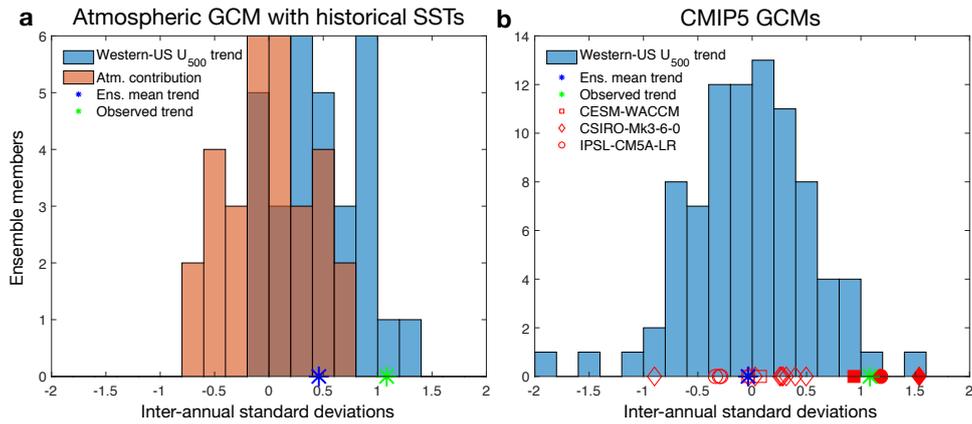
231 The observed circulation trends are likely driven in part by changes in the spatial  
 232 pattern of global sea-surface temperatures (SSTs) (Fig. 3c). In particular, suppressed  
 233 warming in the eastern tropical Pacific and enhanced warming in the northwestern and  
 234 southwestern Pacific are consistent with a shift toward the cool phase of the Interdecadal  
 235 Pacific Oscillation (IPO) (Power et al., 1999) between 1983-84 and 2017-18 (Fig. S6a)—

236 a phenomenon which has previously been linked to the recent global-warming “hiatus”  
237 (Dai et al., 2015). In addition to its global cooling effect, the negative IPO phase is also  
238 associated with enhanced  $U_{500}$  over the western US, consistent with observed trends (Fig.  
239 S6b).

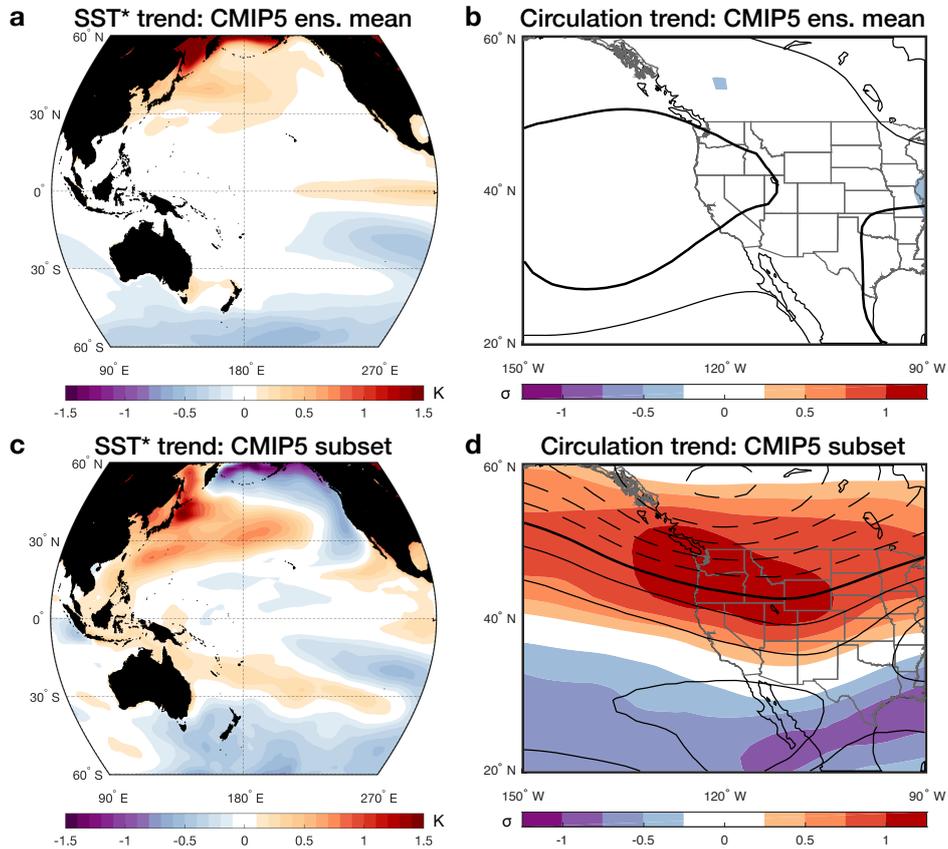
240 To further evaluate the contribution of SSTs to observed circulation trends, we turn  
241 to an ensemble of 30 simulations from the ECHAM5 atmospheric general circulation model  
242 (AGCM) performed using prescribed historical SSTs and radiative forcing. Because the  
243 simulations are identical except for small perturbations in their initial conditions, the  
244 ensemble mean captures the component of atmospheric variability within the simulations  
245 that is driven by changes in SSTs and atmospheric composition, while the ensemble spread  
246 reflects the influence of internal atmospheric variability. Compared with the observed  
247 circulation trends, the ensemble-mean trends have similar spatial structure (Fig. 3d),  
248 but their magnitude is only about half as large over the western US, implying that SSTs  
249 alone do not account for all of the observed circulation trends. On the other hand, ob-  
250 served circulation trends are also not explained by internal atmospheric variability alone,  
251 as indicated by the relatively small ensemble spread in  $U_{500}$  trends over the western US  
252 (Fig. 4a). Given that observed  $U_{500}$  trends fall well within the simulated range (Fig. 4a),  
253 these results indicate that SSTs and internal atmospheric variability have both contributed  
254 to—and together explain—the dynamical suppression of snowpack decline since the 1980s.

## 267 4 Future Implications

268 Looking forward, the implications of our results depend on whether the observed  
269 SST/circulation changes in recent decades have been part of the climate’s response to  
270 anthropogenic forcing, or simply a result of low-frequency natural variability. To answer  
271 this question, we compare the observed SST/circulation trends in Fig. 3b-c with those  
272 simulated by a large ensemble ( $n = 86$ ) of coupled oceanic-atmospheric GCMs from the  
273 latest Coupled Model Intercomparison Project (CMIP5) (Taylor et al., 2012) over the  
274 same 35-year time period (Fig. 5; Table S1). In contrast to the observed trends, the ensemble-  
275 mean of these simulations shows little evidence of an IPO-like shift in SST patterns (Fig.  
276 5a), or of a substantial change in the atmospheric circulation over the western US (Fig.  
277 5b), suggesting that observed trends are not primarily driven by anthropogenic forcing.



255 **Figure 4.** a) The distribution of trends in winter-mean  $U_{500}$  over the western US between  
 256 1983-84 and 2017-18, normalized by the detrended inter-annual standard deviation, within a  
 257 30-member ensemble of simulations from the ECHAM5 atmospheric GCM run with prescribed  
 258 historical SSTs. Blue represents the full distribution, while orange represents the atmospheric  
 259 (i.e., non-SST-driven) component of the distribution, as measured by the departure of each en-  
 260 semble member from the ensemble-mean trend (blue star). The observed trend (green star) falls  
 261 within the full distribution, but outside the atmospheric component of the distribution, indi-  
 262 cating that SST changes likely account for some of the observed trend. b) As in (a), but for an  
 263 86-member ensemble of coupled GCM simulations from CMIP5. Solid red shapes show the three  
 264 largest  $U_{500}$  trends among all ensemble members, which were simulated by the following models:  
 265 CESM-WACCM (square), CSIRO-Mk3-6-0 (diamond), and IPSL-CM5A-LR (circle). Hollow red  
 266 shapes show the trends across all other simulations performed by the same GCMs.



278 **Figure 5.** a) The ensemble-mean trend in winter-mean (October-March) SST\* in CMIP5  
 279 simulations run with historical and projected radiative forcing between 1983-84 and 2017-18. b)  
 280 The ensemble-mean trend in normalized winter-mean  $U_{500}$  (colors) and  $Z_{500}$  (contours). c) As in  
 281 (a), but for a subset of three ensemble members (CESM-WACCM, CSIRO-Mk3-6-0, and IPSL-  
 282 CM5A-LR) that exhibited trends in  $U_{500}$  over the western US that were similar to the observed  
 283 trend. d) As in (b), but for the same subset of simulations included in (c).

284 On the other hand, the observed trend in  $U_{500}$  over the western US does fall within  
285 the range that one would expect from natural variability, based on the CMIP5 ensem-  
286 ble spread (Fig. 4b). Furthermore, among the ensemble members that exhibit the largest  
287  $U_{500}$  trends over the western US (Fig. 5d), SST trends resemble the observed shift to-  
288 ward the cool phase of the IPO (Fig. 5c), supporting our hypothesis that circulation trends  
289 have been driven in part by shifting SST patterns. Meanwhile, the same subset of GCMs  
290 produced a wide range of  $U_{500}$  trends in other identical simulations (Fig. 4b, red shapes),  
291 confirming that the CMIP5 ensemble spread is primarily a reflection of natural variabil-  
292 ity, and not of differences in model physics.

293 Of course, it is possible that observed SST/circulation trends represent a compo-  
294 nent of the forced response that most GCMs do not capture (Kohyama & Hartmann,  
295 2017). If so, then changes in the atmospheric circulation may continue to offset some of  
296 the decline in western-US snowpack due to warming. But if recent circulation changes  
297 have instead been a result of natural variability—the more likely outcome in our view,  
298 given the close correspondence between observed SST changes and the IPO—then the  
299 relative stability of western-US snowpack since the 1980s is unlikely to persist, portend-  
300 ing an accelerated decline in coming decades as the phase of variability becomes less fa-  
301 vorable for snowpack accumulation.

## 302 **Acknowledgments**

303 C.P. was supported by a postdoctoral fellowship from the University of Washington Joint  
304 Institute for the Study of the Atmosphere and the Ocean. The work of S. P. was per-  
305 formed under the auspices of the U.S. Department of Energy (DOE) by LLNL under Con-  
306 tract DE-AC52-07NA27344 with support provided by the LLNL-LDRD Program under  
307 Project 18-ERD-054.

308 Model output and observations can be accessed at the following websites. Snow-  
309 water equivalent and temperature data at each SNOTEL station: <https://www.wcc.nrcs.usda.gov/snow/>. Gridded SST data: <https://www.esrl.noaa.gov/psd/data/gridded/data.noaa.oisst.v2.html>. MERRA-2 reanalysis: <https://gmao.gsfc.nasa.gov/reanalysis/MERRA-2/>. GISTEMP gridded surface temperature analysis: <https://www.esrl.noaa.gov/psd/>. ECHAM5 prescribed-SST simulations: <https://www.esrl.noaa.gov/psd/repository/alias/facts/>. CMIP5 simulations: <https://esgf-node.llnl.gov/search/cmip5/>.

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